

Riemannian geometry for statistical estimation and learning: applications to remote sensing and M/EEG

Antoine Collas

TAU seminar



université
PARIS-SACLAY

Education and Research

- **2022 - Present: Postdoctoral Researcher in Machine Learning**
Mind team (ex-Parietal) at Inria Saclay
Advisors: Alexandre Gramfort, Rémi Flamary
- **2019-2022: PhD in Signal processing**
SONDRA laboratory, CentraleSupélec, University of Paris-Saclay
Directors: Jean-Philippe Ovarlez, Guillaume Ginolhac
- **2014-2019: Engineering degree in Computer Science & Applied Mathematics**
University of Technology of Compiègne (UTC)

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Context

Context in remote sensing

In recent years, many image time series have been taken from the **earth** with different technologies: **SAR, multi/hyper spectral imaging, ...**

Objective

Segment semantically these data using **sensor diversity** (spectral bands, polarization...), and **spatial** and/or **temporal** informations.

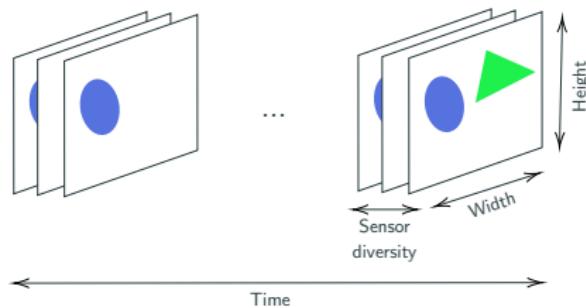


Figure 1: Multivariate image time series.

Applications

Activity monitoring, land cover mapping, crop type mapping, disaster assessment ...

Context in neuroscience

Many new datasets are available in neuroscience: **EEG, MEG, fMRI, ...**

Objectives

- **Classify** brain signals into different **cognitive states** (sleep, wake, anesthesia, seizure, ...).
- **Regress** biomarkers (e.g. age) from brain signals.

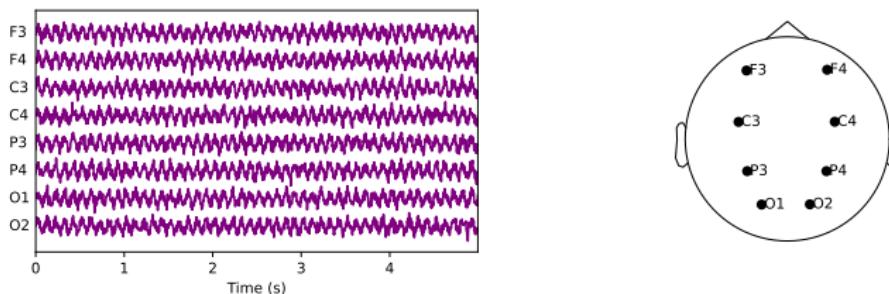
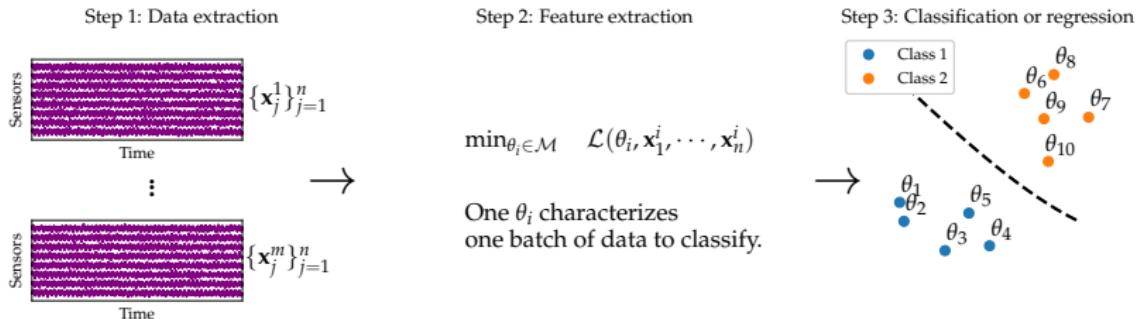


Figure 2: Multivariate EEG time series and the sensor locations.

Applications

Brain-computer interfaces, sleep monitoring, brain aging, ...

Classification and regression pipeline



Assumption:

$\mathbf{x} \sim f(\cdot, \theta)$, a parametric probability density function, $\theta \in \mathcal{M}$

Examples of θ :

$\theta = \Sigma$ a covariance matrix, $\theta = (\mu, \Sigma)$ a vector and a covariance matrix,
 $\theta = (\{\tau_i\}, U)$ a scalar and an orthogonal matrix...

\mathcal{M} can be constrained !

Step 2: objectives for feature estimation



Figure 3: Example of a SAR image (from nasa.gov).

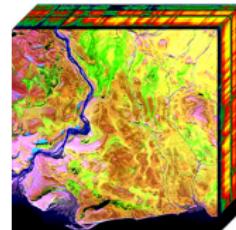


Figure 4: Example of a hyperspectral image (from nasa.gov).

Objectives:

- develop **robust estimators**, *i.e.* estimators for non Gaussian or heterogeneous data because of the high resolution of images and the presence of outliers in biosignals,
- develop **regularized/structured estimators**, *i.e.* estimators that handle the high dimension of hyperspectral images and MEG.

Step 3: objectives for classification and regression

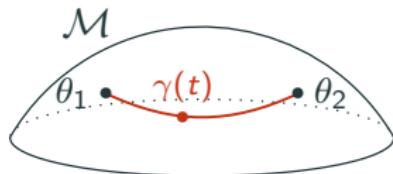


Figure 5: Divergence δ_γ :
squared length of the curve γ .

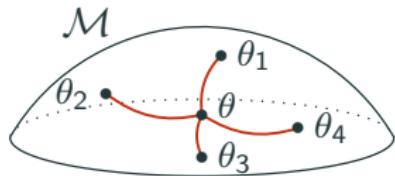


Figure 6: Center of mass of $\{\theta_i\}_{i=1}^M$.

Objectives:

Develop divergences that

- respect the constraints of \mathcal{M} ,
- are related to the chosen statistical distributions,
- are robust to distribution shifts between train and test data.

Use normalizations on \mathcal{M} to fix **distribution shifts** between train and test sets.

Classification and regression pipeline and Riemannian geometry

Random variable: $x \sim f(\cdot; \theta)$, $\theta \in \mathcal{M}$

Step 2: maximum likelihood estimation

$$\underset{\theta \in \mathcal{M}}{\text{minimize}} \mathcal{L}(\theta, \{x_i\}_{i=1}^n) = -\log f(\{x_i\}_{i=1}^n, \theta)$$

Step 3: given δ , center of mass of $\{\theta_i\}_{i=1}^M$

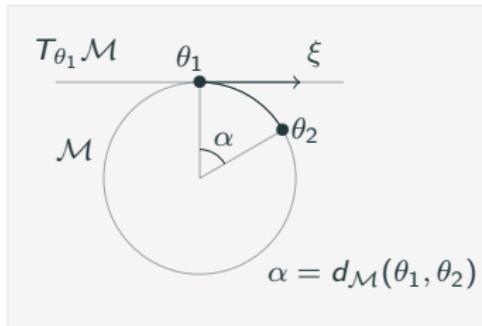
$$\underset{\theta \in \mathcal{M}}{\text{minimize}} \sum_i \delta(\theta, \theta_i)$$

Use of Riemannian geometry:

- optimization under constraints,
- “Fisher information metric” \implies a canonical Riemannian manifold for the parameter space \mathcal{M} (fast estimators, intrinsic Carmér-Rao bounds...),
- δ : squared Riemannian distance.

Riemannian geometry and problematics

What is a Riemannian manifold ?



Curvature induced by:

- constraints, e.g. the sphere: $\|x\| = 1$,
- Riemannian metric, e.g. on S_p^{++} :
$$\langle \xi, \eta \rangle_{\Sigma}^{S_p^{++}} = \text{Tr}(\Sigma^{-1} \xi \Sigma^{-1} \eta).$$

Some geometric tools:

- **tangent space** $T_{\theta}\mathcal{M}$ (vector space): linearization of \mathcal{M} at $\theta \in \mathcal{M}$,
- **Riemannian metric** $\langle ., . \rangle_{\theta}^{\mathcal{M}}$: inner product on $T_{\theta}\mathcal{M}$,
- **geodesic** γ : curve on \mathcal{M} with zero acceleration,
- **distance**: $d_{\mathcal{M}}(\theta_1, \theta_2) = \text{length of } \gamma$.

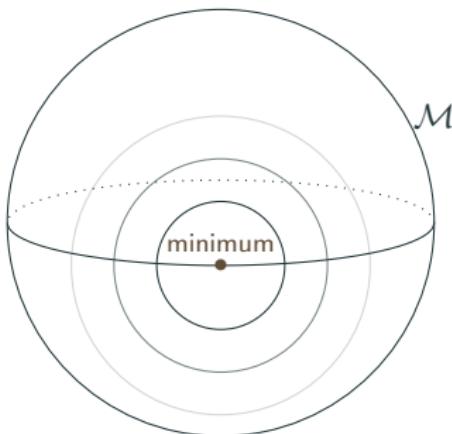
Examples of \mathcal{M} : $\mathbb{R}^{p \times k}$, the sphere S^{p-1} , symmetric positive definite matrices S_p^{++} , orthonormal k -frames $\text{St}_{p,k}$, low-rank matrices, ...

Optimization on a manifold

Optimization

$\mathcal{L} : \mathcal{M} \rightarrow \mathbb{R}$, smooth

$$\underset{\theta \in \mathcal{M}}{\text{minimize}} \mathcal{L}(\theta)$$

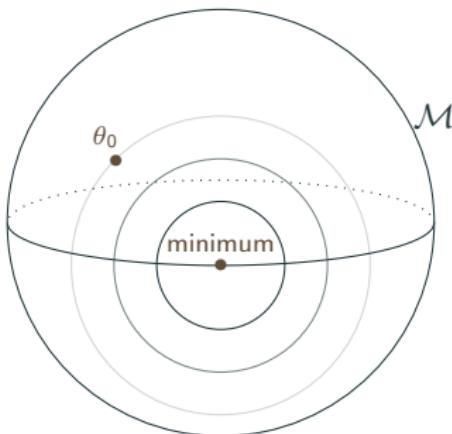


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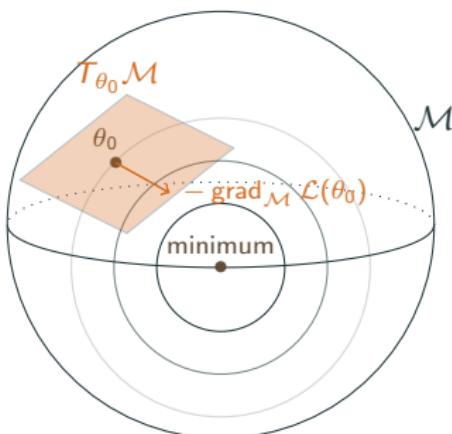


Optimization on a manifold

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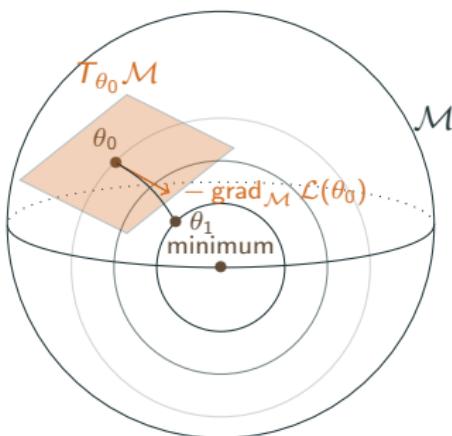


Optimization on a manifold

Optimization

$\mathcal{L} : \mathcal{M} \rightarrow \mathbb{R}$, smooth

$$\underset{\theta \in \mathcal{M}}{\text{minimize}} \mathcal{L}(\theta)$$



Fisher information metric

Random variable, negative log-likelihood

$$\mathbf{x} \sim f(\cdot, \theta), \quad \theta \in \mathcal{M}$$

$$\mathcal{L}(\theta, \mathbf{x}) = -\log f(\mathbf{x}, \theta)$$

Fisher information metric

$$\begin{aligned}\langle \xi, \eta \rangle_{\theta}^{\text{FIM}} &= \mathbb{E}_{\mathbf{x} \sim f(\cdot, \theta)} [D^2 \mathcal{L}(\theta, \mathbf{x}) [\xi, \eta]] \\ &= \text{vec}(\xi)^T I(\theta) \text{vec}(\eta)\end{aligned}$$

where

$$I(\theta) = \mathbb{E}_{\mathbf{x} \sim f(\cdot, \theta)} [\text{Hess } \mathcal{L}(\theta, \mathbf{x})] \in \mathcal{S}_p^{++}$$

is the Fisher information matrix.

(Set of constraints, Fisher information metric) = a Riemannian manifold

Existing work: centered Gaussian

A well known geometry:

$$\mathbf{x} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}), \quad \boldsymbol{\Sigma} \in \mathcal{S}_p^{++}$$

with the Fisher information metric:

$$\langle \boldsymbol{\xi}, \boldsymbol{\eta} \rangle_{\boldsymbol{\Sigma}}^{\text{FIM}} = \text{Tr} (\boldsymbol{\Sigma}^{-1} \boldsymbol{\xi} \boldsymbol{\Sigma}^{-1} \boldsymbol{\eta}).$$

Induced pipeline

Step 2:

$$\hat{\boldsymbol{\Sigma}}_{\text{SCM}} = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i^T.$$

Step 3: geodesic distance on \mathcal{S}_p^{++}

$$d_{\mathcal{S}_p^{++}}(\boldsymbol{\Sigma}_1, \boldsymbol{\Sigma}_2) = \left\| \log \left(\boldsymbol{\Sigma}_1^{-\frac{1}{2}} \boldsymbol{\Sigma}_2 \boldsymbol{\Sigma}_1^{-\frac{1}{2}} \right) \right\|_2.$$

Riemannian gradient descent to solve:

$$\underset{\boldsymbol{\Sigma} \in \mathcal{S}_p^{++}}{\text{minimize}} \sum_i d_{\mathcal{S}_p^{++}}^2(\boldsymbol{\Sigma}, \boldsymbol{\Sigma}_i).$$

Alexandre Barachant et al. "Multiclass Brain–Computer Interface Classification by Riemannian Geometry". In: *IEEE Transactions on Biomedical Engineering* 59.4 (2012), pp. 920–928

Problematics

Go beyond $x \sim \mathcal{N}(\mathbf{0}, \Sigma)$

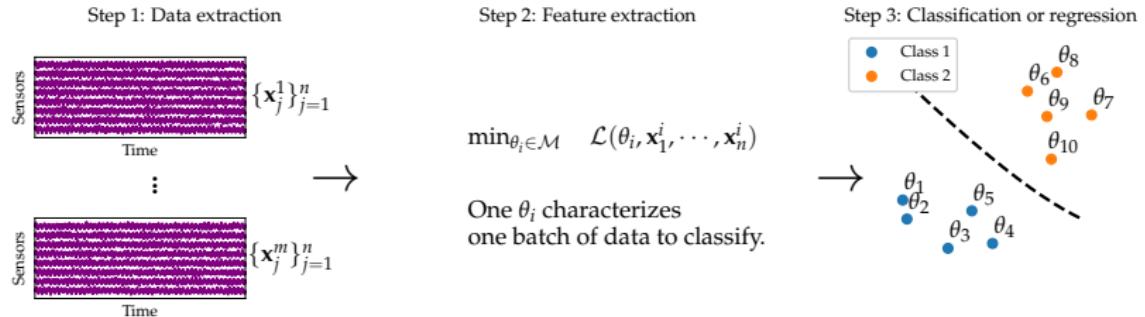
- $x_i \sim \mathcal{N}(\mu, \tau_i \Sigma)$ for non-centered data and robustness,
- $x_i \sim \mathcal{N}(\mathbf{0}, \tau_i \mathbf{U} \mathbf{U}^T + \mathbf{I}_p)$ for high dimensional data and robustness.

Problems

- Existence of maximum likelihood estimators ?
- Not always closed form estimators: how to get fast iterative algo. ?
- Not always closed form expression of the Riemannian distance: what to do ?
- How to get fast estimators of centers of mass ?

Estimation and classification of non centered and heteroscedastic data

Non-centered mixtures of scaled Gaussian distributions



Non-centered mixtures of scaled Gaussian distributions (NC-MSGs)

Let $x_1, \dots, x_n \in \mathbb{R}^p$ distributed as $x_i \sim \mathcal{N}(\mu, \tau_i \Sigma)$ with $\mu \in \mathbb{R}^p$, $\Sigma \in S_p^{++}$, and $\tau \in (\mathbb{R}_*^+)^n$.

Goal: estimate and classify $\theta = (\mu, \Sigma, \tau)$.

Interesting when data are heteroscedastic (e.g. time series) and/or contain outliers.

Parameter space and cost functions

Parameter space: location, scatter matrix, and textures

$$\mathcal{M}_{p,n} = \mathbb{R}^p \times \mathcal{S}_p^{++} \times \mathcal{S}(\mathbb{R}_*^+)^n$$

where

$$\mathcal{S}(\mathbb{R}_*^+)^n = \left\{ \boldsymbol{\tau} \in (\mathbb{R}_*^+)^n : \prod_{i=1}^n \tau_i = 1 \right\}$$

- Positivity constraints: $\boldsymbol{\Sigma} \succ \mathbf{0}$, $\tau_i > 0$
- Scale constraint: $\prod_{i=1}^n \tau_i = 1$

Parameter estimation

Minimization of a regularized negative log-likelihood (NLL), $\beta \geq 0$

$$\underset{\theta \in \mathcal{M}_{p,n}}{\text{minimize}} \quad \mathcal{L}(\theta, \{\mathbf{x}_i\}_{i=1}^n) + \beta \mathcal{R}_\kappa(\theta)$$

Center of mass estimation

Averaging parameters $\{\theta_i\}_{i=1}^M$ with a to be defined divergence δ

$$\underset{\theta \in \mathcal{M}_{p,n}}{\text{minimize}} \quad \frac{1}{M} \sum_{i=1}^M \delta(\theta, \theta_i)$$

Parameter space with a product metric

Product metric

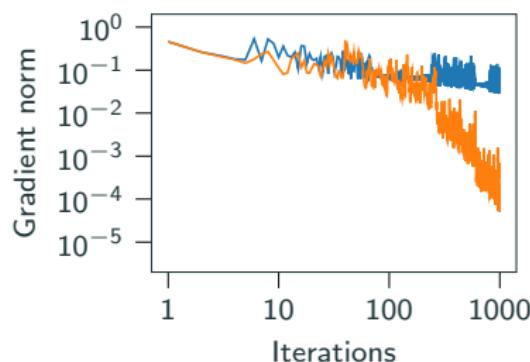
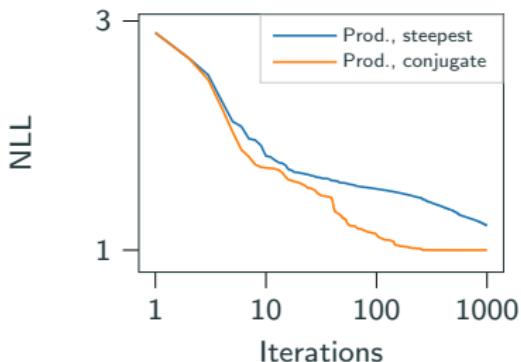
Let $\xi = (\xi_\mu, \xi_\Sigma, \xi_\tau)$, $\eta = (\eta_\mu, \eta_\Sigma, \eta_\tau)$ in the tangent space,

$$\langle \xi, \eta \rangle_{\theta}^{\mathcal{M}_{p,n}^{\text{Prod.}}} = \xi_\mu^T \eta_\mu + \text{Tr}(\Sigma^{-1} \xi_\Sigma \Sigma^{-1} \eta_\Sigma) + (\xi_\tau \odot \tau^{\odot -1})^T (\eta_\tau \odot \tau^{\odot -1})$$

where \odot is the elementwise operator.

Product manifold \implies Riemannian conjugate gradient on $(\mathcal{M}_{p,n}, \langle ., . \rangle^{\mathcal{M}_{p,n}^{\text{Prod.}}})$.

Slow in practice ...



Parameter space with the Fisher information metric

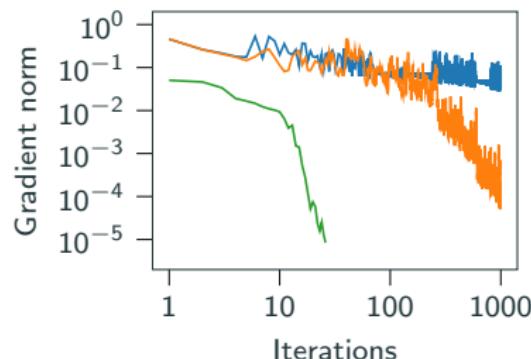
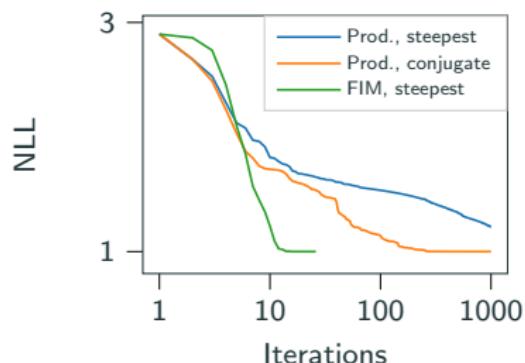
Fisher information metric of NC-MSGs

Let $\xi = (\xi_\mu, \xi_\Sigma, \xi_\tau)$, $\eta = (\eta_\mu, \eta_\Sigma, \eta_\tau)$ in the tangent space,

$$\langle \xi, \eta \rangle_{\theta}^{\mathcal{M}_{p,n}^{\text{FIM}}} = \sum_{i=1}^n \frac{1}{\tau_i} \xi_\mu^T \Sigma^{-1} \eta_\mu + \frac{n}{2} \text{Tr}(\Sigma^{-1} \xi_\Sigma \Sigma^{-1} \eta_\Sigma) + \frac{p}{2} (\xi_\tau \odot \tau^{\odot -1})^T (\eta_\tau \odot \tau^{\odot -1})$$

Derivation of the Riemannian gradient and a second order retraction.

\Rightarrow Riemannian gradient descent on $(\mathcal{M}_{p,n}, \langle ., . \rangle_{\cdot}^{\mathcal{M}_{p,n}^{\text{FIM}}})$.



Parameter estimation: existence

Observation of sequences $(\theta^{(\ell)})_\ell$ such that

$$\mathcal{L}(\theta^{(\ell+1)}) < \mathcal{L}(\theta^{(\ell)}) \quad \text{and} \quad \theta^{(\ell)} \xrightarrow[\ell \rightarrow +\infty]{} \partial\theta$$

where $\partial\theta$ is a border of $\mathcal{M}_{p,n}$ (e.g. $\tau_i = 0$).

Existence of a regularized maximum likelihood estimator

Under some assumptions on \mathcal{R}_κ and $\beta > 0$, the regularized NLL

$$\theta \mapsto \mathcal{L}(\theta, \{\mathbf{x}_i\}_{i=1}^n) + \beta \mathcal{R}_\kappa(\theta),$$

admits a minimum in $\mathcal{M}_{p,n}$.

Example:

$$\mathcal{R}_\kappa(\theta) = \sum_{i,j} \left((\tau_i \lambda_j)^{-1} - \kappa^{-1} \right)^2$$

where λ_j are the eigenvalues of $\boldsymbol{\Sigma}$.

Classification

KL divergence between NC-MSGs

$$\delta_{\text{KL}}(\theta_1, \theta_2) \propto \sum_{i=1}^n \frac{\tau_{1,i}}{\tau_{2,i}} \text{Tr} \left(\boldsymbol{\Sigma}_2^{-1} \boldsymbol{\Sigma}_1 \right) + \sum_{i=1}^n \frac{1}{\tau_{2,i}} \Delta \boldsymbol{\mu}^T \boldsymbol{\Sigma}_2^{-1} \Delta \boldsymbol{\mu} + n \log \left(\frac{|\boldsymbol{\Sigma}_2|}{|\boldsymbol{\Sigma}_1|} \right)$$

with $\Delta \boldsymbol{\mu} = \boldsymbol{\mu}_2 - \boldsymbol{\mu}_1$.

Symmetrization: $\delta_{\mathcal{M}_{p,n}}(\theta_1, \theta_2) = \frac{1}{2} (\delta_{\text{KL}}(\theta_1, \theta_2) + \delta_{\text{KL}}(\theta_2, \theta_1))$.

Riemannian center of mass

Minimization of the KL variance:

$$\underset{\theta \in \mathcal{M}_{p,n}}{\text{minimize}} \frac{1}{M} \sum_{i=1}^M \delta_{\mathcal{M}_{p,n}}(\theta, \theta_i)$$

Done with a Riemannian gradient descent.

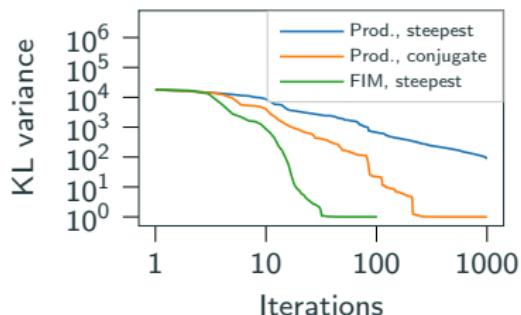


Figure 9: KL variance vs. iterations with $p = 10$, $n = 150$ and $M = 2$.

Breizhcrops dataset

Breizhcrops dataset¹:

- more than 600 000 crop time series across the whole Brittany,
- 13 spectral bands, 9 classes.

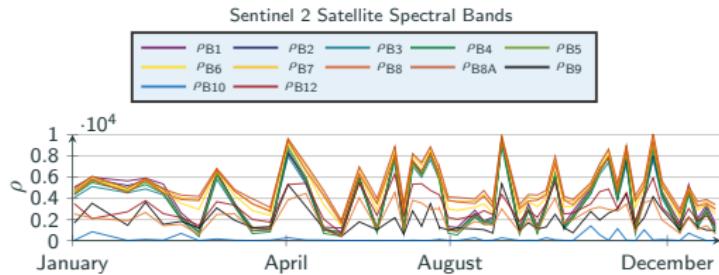


Figure 10: Reflectances ρ of a time series of **meadows**.

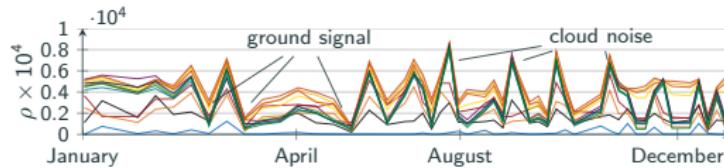


Figure 11: Reflectances ρ of a time series of **corn**.

¹<https://breizhcrops.org/>

Application to the Breizhcrops dataset

Parameter estimation + classification with a *Nearest centroid classifier*

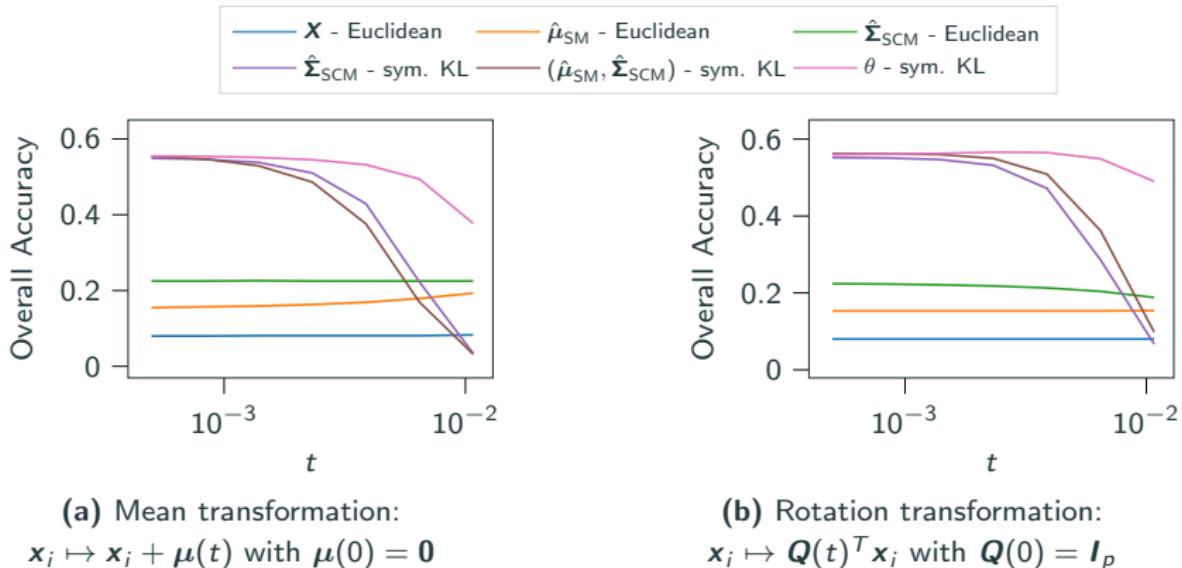
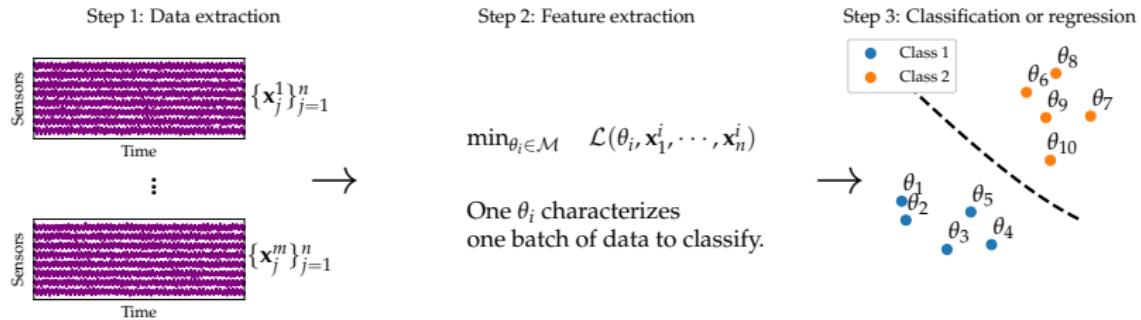


Figure 12: “Overall Accuracy” metric versus the parameter t associated with transformations applied to the test set. The proposed *Nearest centroid classifier* is “ θ - sym. KL”. The regularization is the L2 penalty and $\beta = 10^{-11}$.

Probabilistic PCA from heteroscedastic signals

Study of a “low rank” statistical model



Statistical model

$\mathbf{x}_1, \dots, \mathbf{x}_n \in \mathbb{R}^p, \forall k < p:$

$$\mathbf{x}_i \sim \mathcal{N}(\mathbf{0}, \tau_i \mathbf{U} \mathbf{U}^T + \mathbf{I}_p)$$

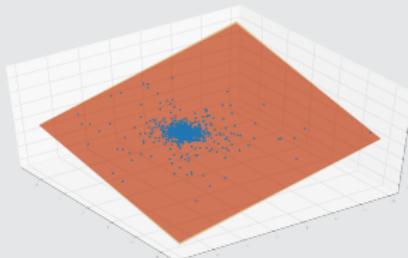
with $\tau_i > 0$ and $\mathbf{U} \in \mathbb{R}^{p \times k}$ is an orthogonal basis ($\mathbf{U}^T \mathbf{U} = \mathbf{I}_k$).

Goal: estimate and classify $\theta = (\mathbf{U}, \tau)$.

Study of a “low rank” statistical model

Statistical model

$$\underbrace{\mathbf{x}_i}_{\in \mathbb{R}^p} \stackrel{d}{=} \underbrace{\sqrt{\tau_i} \mathbf{U} \mathbf{g}}_{\text{signal} \in \text{span}(\mathbf{U})} + \underbrace{\mathbf{n}}_{\text{noise} \in \mathbb{R}^p}$$



where $\mathbf{g} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_k) \perp \mathbf{n} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$, $\tau \in (\mathbb{R}_*^+)^n$, and $\mathbf{U} \in \mathbb{R}^{p \times k}$ s.t. $\mathbf{U}^T \mathbf{U} = \mathbf{I}_k$.

Maximum likelihood estimation

Minimization of the NLL with constraints, $\theta = (\mathbf{U}, \tau)$

- $\mathbf{U} \in \text{Gr}_{p,k}$: orthogonal basis of the subspace (and thus invariant by rotation !)
- $\tau \in (\mathbb{R}_*^+)^n$: positivity constraints

$$\underset{\theta \in \text{Gr}_{p,k} \times (\mathbb{R}_*^+)^n}{\text{minimize}} \mathcal{L}(\theta, \{\mathbf{x}_i\}_{i=1}^n)$$

Study of a “low rank” statistical model: estimation

Fisher information metric

$\forall \xi = (\xi_U, \xi_\tau), \eta = (\eta_U, \eta_\tau)$ in the tangent space

$$\langle \xi, \eta \rangle_{\theta}^{\text{FIM}} = 2nc_\tau \operatorname{Tr} \left(\xi_U^T \eta_U \right) + k (\xi_\tau \odot (1 + \tau)^{\odot -1})^T (\eta_\tau \odot (1 + \tau)^{\odot -1}),$$

$$\text{where } c_\tau = \frac{1}{n} \sum_{i=1}^n \frac{\tau_i^2}{1 + \tau_i}.$$

Derivation of the Riemannian gradient and of a retraction.

To minimize the NLL: Riemannian gradient descent on $(\text{Gr}_{p,k} \times (\mathbb{R}_*)^n, \langle \cdot, \cdot \rangle^{\text{FIM}})$.

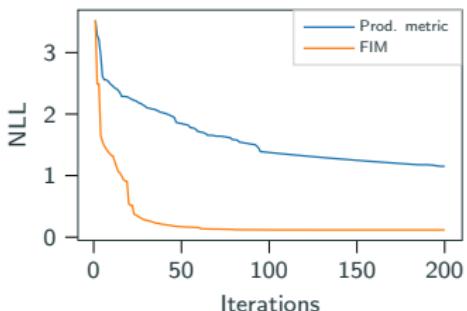


Figure 13: NLL versus the iterations.

Study of a “low rank” statistical model: bounds

Intrinsic Cramér-Rao bounds

Study of the performance through intrinsic Cramér-Rao bounds:

$$\underbrace{\mathbb{E}[d_{\text{Gr}_{p,k}}^2(\text{span}(\hat{\mathbf{U}}), \text{span}(\mathbf{U}))]}_{\text{subspace estimation error}} \geq \frac{(p-k)k}{nc_\tau} \approx \frac{(p-k)k}{n \times \text{SNR}}$$
$$\underbrace{\mathbb{E}[d_{(\mathbb{R}_*^+)^n}^2(\hat{\boldsymbol{\tau}}, \boldsymbol{\tau})]}_{\text{texture estimation error}} \geq \frac{1}{k} \sum_{i=1}^n \frac{(1 + \tau_i)^2}{\tau_i^2}$$

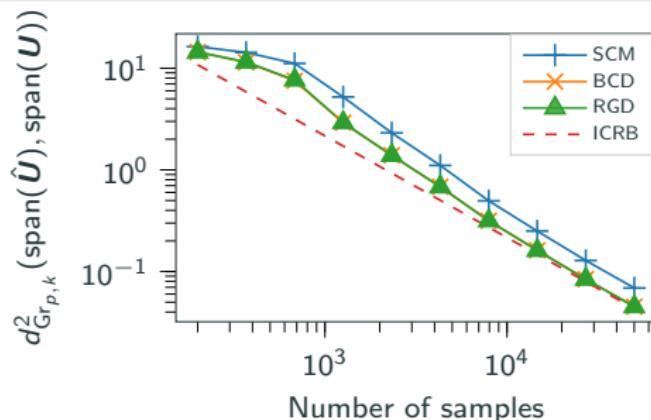


Figure 14: Mean squared error versus the number of simulated data.

Aligning M/EEG data to enhance predictive regression modeling

Generative model for regression with M/EEG

Linear instantaneous mixing model (from Maxwell's equations)

Signal $\mathbf{h}(t) \sim \mathcal{N}(\mathbf{0}, \Sigma)$:

$$\underbrace{\mathbf{h}(t)}_{\text{observed signal}} = \mathbf{A} \underbrace{\boldsymbol{\eta}(t)}_{\text{latent sources}}$$

Covariance matrix:

$$\Sigma = \mathbb{E}_t [\mathbf{h}(t)\mathbf{h}(t)] = \mathbf{A} \text{diag}(\mathbf{p}) \mathbf{A}^\top$$

with $\mathbf{p} = \text{Var}(\boldsymbol{\eta}(t))$.

Regression model, $(\Sigma_i, y_i)_{i=1}^m$

If $\exists \beta \in \mathbb{R}^p$ s.t.

$$y_i = \beta^\top \log(\mathbf{p}_i) + \varepsilon_i$$

then $\exists \beta' \in \mathbb{R}^{p(p+1)/2}$ s.t.

$$y_i = \beta'^\top \underbrace{\text{vec} \left(\log(\bar{\Sigma}^{-\frac{1}{2}} \Sigma_i \bar{\Sigma}^{-\frac{1}{2}}) \right)}_{\in T_I S_p^{++}} + \varepsilon_i$$

where $\bar{\Sigma}$ is the Riemannian mean of $\{\Sigma_i\}_{i=1}^m$.

David Sabbagh et al. "Manifold-regression to predict from MEG/EEG brain signals without source modeling". In: *Advances in Neural Information Processing Systems 32* (2019)

A. Mellot, A. Collas et al. "Harmonizing and aligning M/EEG datasets with covariance-based techniques to enhance predictive regression modeling" in *Imaging Neuroscience* MIT Press 2023

Statistics on the \mathcal{S}_p^{++} manifold

Gaussian distribution on \mathcal{S}_p^{++} and normalization

$$f(\boldsymbol{\Sigma}; \bar{\boldsymbol{\Sigma}}, \sigma^2) = \frac{1}{Z(\sigma)} \exp\left(-\frac{d_{\mathcal{S}_p^{++}}^2(\boldsymbol{\Sigma}, \bar{\boldsymbol{\Sigma}})}{2\sigma^2}\right)$$

with $Z(\sigma)$ the normalization constant.

Recenter-rescale operator: $\phi_{\bar{\boldsymbol{\Sigma}}, \sigma^2}(\boldsymbol{\Sigma}) = \left(\bar{\boldsymbol{\Sigma}}^{-\frac{1}{2}} \boldsymbol{\Sigma} \bar{\boldsymbol{\Sigma}}^{-\frac{1}{2}}\right)^{\frac{1}{\sigma}}$.

Salem Said et al. "Riemannian Gaussian Distributions on the Space of Symmetric Positive Definite Matrices". In: *IEEE Transactions on Information Theory* 63.4 (2017), pp. 2153–2170

Estimation with $(\boldsymbol{\Sigma}_i)_{i=1}^n \sim f(\cdot; \bar{\boldsymbol{\Sigma}}, \sigma^2)$

$$\hat{\bar{\boldsymbol{\Sigma}}} = \arg \min_{\boldsymbol{\Sigma} \in \mathcal{S}_p^{++}} \frac{1}{n} \sum_{i=1}^n d_{\mathcal{S}_p^{++}}^2(\boldsymbol{\Sigma}, \boldsymbol{\Sigma}_i), \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n d_{\mathcal{S}_p^{++}}^2(\hat{\bar{\boldsymbol{\Sigma}}}, \boldsymbol{\Sigma}_i)$$

Domain adaptation: for $\mathcal{D} \in \{\mathcal{S}, \mathcal{T}\}$

$$\boldsymbol{\Sigma}_i^{\mathcal{D}} \leftarrow \phi_{\hat{\bar{\boldsymbol{\Sigma}}}^{\mathcal{D}}, (\hat{\sigma}^2)^{\mathcal{D}}}(\boldsymbol{\Sigma}_i^{\mathcal{D}})$$

Results on MEG data

Brain age prediction on the Cam-CAN dataset:

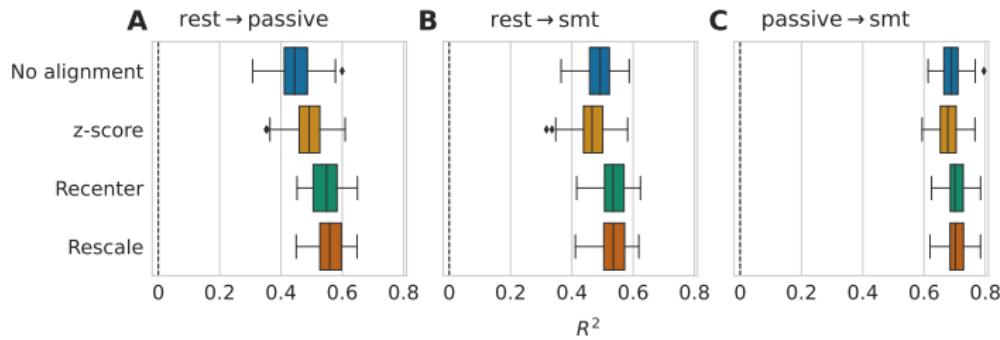


Figure 15: R^2 score on the Cam-CAN dataset (MEG), $n = 646$, 306 channels reduced to $p = 65$ after PCA and age range of 18 – 89 years old.

Results on EEG datasets: LEMON → TUAB

Brain age prediction on the LEMON → TUAB datasets, regression on supervised SPoC components: $\text{diag}(\log(\mathbf{W}_{\text{SPoC}} \boldsymbol{\Sigma}_i \mathbf{W}_{\text{SPoC}}^\top))$.

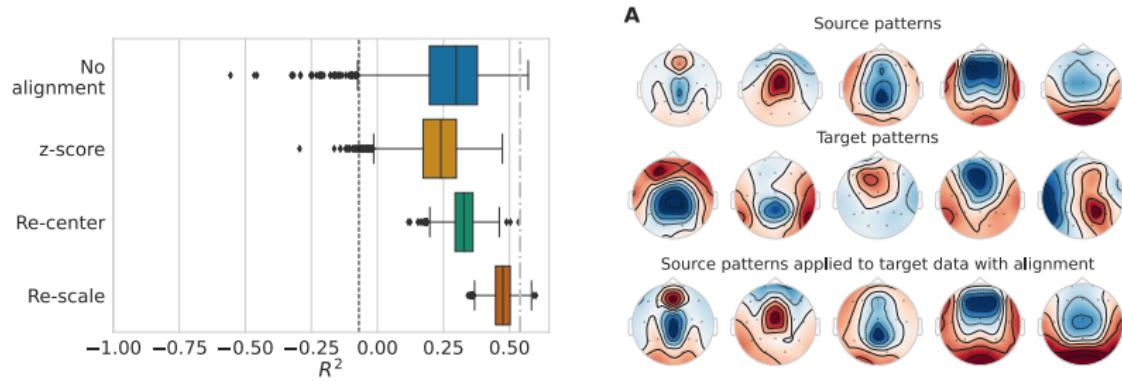
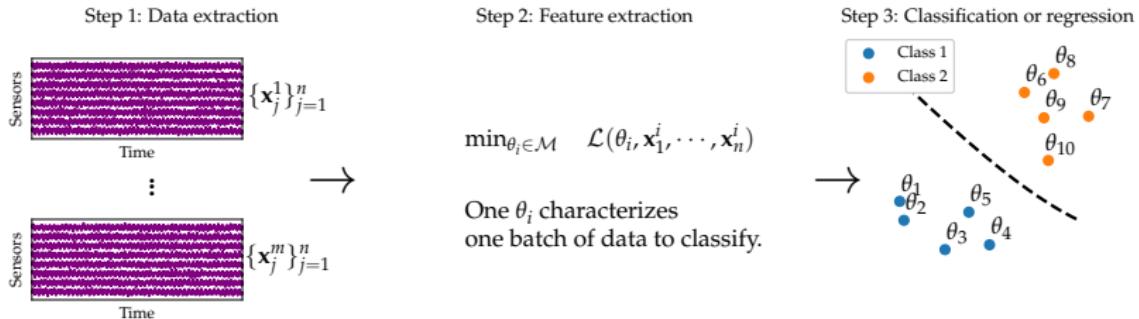


Figure 16: Left: R^2 score on LEMON ($n = 1385$) → TUAB ($n = 213$) (EEG), and $p = 15$ after PCA. Dashed line is the R^2 score of a cross-validation on target dataset. Right: topomaps of the SPoC patterns.

Many other results in the paper: simulations, rotation corrections, ...

Open source software and conclusions

Open source software



pyCovariance (creator): github.com/antoinecollas/pyCovariance

- `_FeatureArray`: custom data structure to store batch of points of product manifolds,
- implements statistical manifolds from this presentation,
- automatic computation of Riemannian centers of mass using `exp`/`log` or `autodiff`
- *K-means++* and *Nearest centroid classifier* on any Riemannian manifolds,
- 15K lines of code, 96% of test coverage.

Open source software

pyManopt (maintainer): github.com/pymanopt/pymanopt

$$\underset{\theta \in \mathcal{M}}{\text{minimize}} f(\theta)$$

Provide f smooth, choose a Riemannian manifold \mathcal{M} , and pyManopt does the rest !

Geomstats: information geometry module (co-creator)
github.com/geomstats/geomstats

Choose a statistical manifold \mathcal{M} (or give a p.d.f. !), and Geomstats does the rest: geodesics, log, exp, barycenter, leaning: K-means, KNN, PCA, etc...

A. Le Brigant, J. Deschamps, **A. Collas** and N. Miolane, “Parametric information geometry with the package Geomstats” ACM Transactions on Mathematical Software 2023.

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Riemannian geometry for statistical estimation and learning: applications to remote sensing and M/EEG

Antoine Collas

TAU seminar



université
PARIS-SACLAY